

1. (20%) 解微分方程式

(a) (10%) $y'(t) = (\tan t)(y^2 - 1)$, $y(0) = 2$. 求 $y(t)$, 其中 t 在 0 的附近。(b) (10%) $t^2 y'(t) - 2ty(t) = 3t^{\frac{3}{2}}$, $y(1) = 0$. 求 $y(t)$, 其中 $t > 0$.**Solution:**(a) 注意到 $y(0) = 2$, $y(t)$ 的連續性保證在 $t = 0$ 的附近, $y(t) > 1$, 使用分離變數:

$$\begin{aligned}
y'(t) &= (\tan t)(y^2(t) - 1) \\
\Rightarrow \frac{y'(t)}{y^2(t) - 1} &= \tan(t) && (y^2(t) - 1 \neq 0) \\
\Rightarrow \int_0^t \frac{y'(s)}{y^2(s) - 1} ds &= \int_0^t \tan(s) ds \\
\Rightarrow \frac{1}{2} \int_0^t \frac{dy(s)}{y(s) - 1} - \frac{1}{2} \int_0^t \frac{dy(s)}{y(s) + 1} &= [-\ln |\cos(s)|]_{s=0}^{s=t} \\
\Rightarrow \frac{1}{2} \left[\ln \left| \frac{y(s) - 1}{y(s) + 1} \right| \right]_{s=0}^{s=t} &= -\ln |\cos(t)| \\
\Rightarrow \frac{1}{2} \ln \left| \frac{y(t) - 1}{y(t) + 1} \right| - \frac{1}{2} \ln \left| \frac{y(0) - 1}{y(0) + 1} \right| &= -\ln |\cos(t)| \\
\Rightarrow \frac{1}{2} \ln \left| \frac{y(t) - 1}{y(t) + 1} \right| - \frac{1}{2} \ln \left(\frac{1}{3} \right) &= -\ln |\cos(t)| && (y(0) = 2) \\
\Rightarrow \ln \left| \frac{y(t) - 1}{y(t) + 1} \right| &= \ln \left(\frac{1}{3} \right) - 2 \ln |\cos(t)| \\
\Rightarrow \left| \frac{y(t) - 1}{y(t) + 1} \right| &= \frac{1}{3 \cos^2(t)} \\
\Rightarrow \frac{y(t) - 1}{y(t) + 1} &= \frac{1}{3 \cos^2(t)} && (y(t) > 1) \\
\Rightarrow y(t) &= \frac{3 \cos^2(t) + 1}{3 \cos^2(t) - 1}
\end{aligned}$$

配分:

- (3%) 分離變數並積分
- (4%) 積分正確
- (2%) 初始條件代入正確
- (1%) 最終答案

(b) 這是一階線性常微分方程，利用積分因子：

$$\begin{aligned}t^2 y'(t) - 2ty(t) &= 3t^{\frac{3}{2}} && (t \neq 0) \\ \Rightarrow y'(t) - 2t^{-1}y(t) &= 3t^{-\frac{1}{2}} \\ \Rightarrow t^{-2} \cdot (y'(t) - 2t^{-1}y(t)) &= t^{-2} \cdot 3t^{-\frac{1}{2}} \quad (\text{積分因子為 } u(t) = \exp\left(\int -2t^{-1}dt\right) = t^{-2}) \\ \Rightarrow (t^{-2}y(t))' &= 3t^{-\frac{5}{2}} \\ \Rightarrow \int_1^t (s^{-2}y(s))' ds &= \int_1^t 3s^{-\frac{5}{2}} ds \\ \Rightarrow \left[\frac{y(s)}{s^2}\right]_{s=1}^{s=t} &= \left[-2s^{-\frac{3}{2}}\right]_{s=1}^{s=t} \\ \Rightarrow \frac{y(t)}{t^2} - \frac{y(1)}{1^2} &= -2t^{-\frac{3}{2}} + 2 \cdot 1^{-\frac{3}{2}} \\ \Rightarrow \frac{y(t)}{t^2} &= -2t^{-\frac{3}{2}} + 2 && (y(1) = 0) \\ \Rightarrow y(t) &= -2\sqrt{t} + 2t^2\end{aligned}$$

配分：

- (3%) 積分因子正確
- (4%) 積分正確
- (2%) 初始條件代入正確
- (1%) 最終答案

2. (18%) 假設實驗室容器中的果蠅數目在第 t 天時為 $P(t)$ ，且 $P(t)$ 滿足微分方程式 $P'(t) = \lambda P(t)(M - P(t))$ ，其中 $\lambda, M > 0$ ， M 是果蠅數目的上限。

(a) (8%) 已知 $t = t_0 > 0$ 時， $P(t_0) = P_0$ 且 $0 < P_0 < M$ 。求 $P(t)$ 。(請寫出解微分方程式的過程。)

(b) (4%) 若 $P(t) = \tilde{P}$ 時，果蠅數目增加率($P'(t)$)達到最大值，求 \tilde{P} (以 M 表示)。

(c) (6%) 研究人員發現當果蠅數達到 50 隻時，果蠅數目增加速率最大，並且 2 天後果蠅由 50 隻增加到 70 隻。估計再經過 2 天，牠們將由 70 隻增加到幾隻?(四捨五入到整數)

Solution:

(a)

$$P'(t) = \lambda P(t)(M - P(t))$$

$$\frac{P'(t)}{P(t)(M - P(t))} = \lambda$$

$$\frac{P'(t)}{M} \left(\frac{1}{P(t)} + \frac{1}{M - P(t)} \right) = \lambda$$

$$\frac{P'(t)}{1} \left(\frac{1}{P(t)} + \frac{1}{M - P(t)} \right) = \lambda M$$

$$\ln \left(\frac{P(t)}{M - P(t)} \right) \Big|_{P_0}^P = \lambda M(t - t_0)$$

$$\ln \left(\frac{P(t)}{M - P(t)} \right) - \ln \left(\frac{P_0(t)}{M - P_0(t)} \right) = \lambda M(t - t_0)$$

$$\frac{P(t)}{M - P(t)} = e^{\lambda M(t - t_0)} * \left(\frac{P_0(t)}{M - P_0(t)} \right)$$

$$\frac{M - P(t)}{P(t)} = \frac{1}{e^{\lambda M(t - t_0)} * \left(\frac{P_0(t)}{M - P_0(t)} \right)}$$

$$\frac{M}{P(t)} + 1 = \frac{1}{e^{\lambda M(t - t_0)} * \left(\frac{P_0(t)}{M - P_0(t)} \right)}$$

$$P(t) = \frac{M}{\frac{1}{e^{\lambda M(t - t_0)} * \left(\frac{P_0(t)}{M - P_0(t)} \right)} - 1}$$

什麼都沒寫+0

積分成功+6

寫出 I.C +8

(b)

$$P''(t) = (\lambda P(t)M - \lambda P^2(t))' = 0$$

$$(\lambda P(t)M - \lambda P^2(t))' = \lambda P'(t)M - 2P(t)\lambda * P'(t) = 0$$

$$P(t) = \frac{M}{2}$$

直接寫答案+3

有過程算錯+3

(c)

$$P(t) = \frac{M}{\frac{1}{e^{\lambda M(t-t_0)} * \left(\frac{P_0(t)}{M - P_0(t)}\right)} - 1}$$

$$P(t_0) = \frac{M}{2}$$

$$P(t_0 + 2) = 70$$

$$P(t_0 + 4) = 84$$

有想出其他近似方法+4(例如 $\frac{70}{50} = \dots$ ，如果給到+3再來要分~)

若不是這樣算但是算對+6

3. (10%) 假設 X 是一個隨機變數且

$$P_X(-1) = \frac{1}{5}, \quad P_X(0) = \frac{1}{5}, \quad P_X(1) = \frac{2}{5}, \quad P_X(2) = \frac{1}{5}.$$

令 Y 是一個隨機變數滿足 $Y = X^2 + 3$ 。

(a) (4%) 求 Y 的可能取值範圍和它的機率函數。

(b) (6%) 求出 $E(Y)$ 和 $\text{Var}(Y)$ 。

Solution:

(a)

$$Y = X^2 + 3 \Rightarrow Y = 3, 4, 7(2pt)$$

$$f_Y(y = 3) = f_X(x = 0) = \frac{1}{5}$$

$$f_Y(y = 4) = f_X(x = -1) + f_X(x = 1) = \frac{3}{5}$$

$$f_Y(y = 7) = f_X(x = 2) = \frac{1}{5}(2pt)$$

(b)

$$E[Y] = 3 * \frac{1}{5} + 4 * \frac{3}{5} + 7 * \frac{1}{5} = \frac{3 + 12 + 7}{5} = 4.4(2pt)$$

$$E[Y^2] = 3^2 * \frac{1}{5} + 4^2 * \frac{3}{5} + 7^2 * \frac{1}{5} = \frac{9 + 48 + 49}{5} = 21.2(2pt)$$

$$\text{Var}(Y) = E[Y^2] - (E[Y])^2 = \frac{106}{5} - \left(\frac{22}{5}\right)^2 = 1.84(2pt)$$

4. (14%) 丟一個公正的銅板 100 次。令 Z 是出現正面的次數。
- (a) (6%) 求 $E(Z)$ 和 $\text{Var}(Z)$ 。
- (b) (4%) 敘述 Chebyshev 不等式。
- (c) (4%) 用 Chebyshev 不等式估計 $P(35 \leq Z \leq 65)$ 至少是多少。

Solution:

(a) Z 為二項分配 $B(100, \frac{1}{2})$ 。故

$$E(Z) = 100 \times \frac{1}{2} = 50,$$
$$\text{Var}(Z) = 100 \times \frac{1}{2} \times (1 - \frac{1}{2}) = 25.$$

(b) 令 X 為一隨機變數，則 Chebyshev 不等式為

$$P(|X - E(X)| \geq k\sqrt{\text{Var}(X)}) \leq \frac{1}{k^2},$$

或可以寫成為

$$P(|X - E(X)| \geq \varepsilon) \leq \frac{\text{Var}(X)}{\varepsilon^2}.$$

(c) 由前兩小題可以知道

$$\begin{aligned} P(35 \leq Z \leq 65) &= P(|Z - 50| \leq 15) \\ &= 1 - P(|Z - 50| > 15) \\ &\geq 1 - \frac{1}{9} \\ &= \frac{8}{9}. \end{aligned}$$

5. (10%) 設 X 的機率密度函數是 $f_X(t) = \frac{1}{\sqrt{\pi}}e^{-t^2}$ 。令 $Z = X^2$ 。

(a) (4%) 寫下 Z 的分配函數(不需要算出積分)。

(b) (6%) 求 Z 的機率密度函數。

Solution:

(a) CDF (Cumulative Distribution Function) = $F_z(t)$

$$F_z(t) = P(Z \leq t) = P(X^2 \leq t) = P(-\sqrt{t} \leq X \leq \sqrt{t}) = \int_{-\sqrt{t}}^{\sqrt{t}} f_x(s) ds = \int_{-\sqrt{t}}^{\sqrt{t}} \frac{1}{\sqrt{\pi}} e^{-s^2} ds$$

note 1: you better use dummy variable for the integration. If you didn't do it this time, you won't lose any point.

note 2: If you didn't write the process precisely, you will only get 2 points at most.

(b) PDF (Probability density function) = $f_z(t)$

$$f_z(t) = \frac{dF_z}{dt} = \frac{d(\int_{-\sqrt{t}}^{\sqrt{t}} \frac{1}{\sqrt{\pi}} e^{-s^2} ds)}{dt} = 2 \times \frac{d(\int_0^{\sqrt{t}} \frac{1}{\sqrt{\pi}} e^{-s^2} ds)}{dt} = 2 \times \frac{1}{\sqrt{\pi}} e^{-\sqrt{t}^2} \times (\sqrt{t})'$$

(by fundamental thm of calculus)

$$= 2 \times \frac{1}{\sqrt{\pi}} e^{-t} \times \frac{1}{2} \times t^{-\frac{1}{2}} = \frac{1}{\sqrt{\pi}} \times e^{-t} \times t^{-\frac{1}{2}}, t > 0$$

$$0, t \leq 0$$

note 1: you would get 3 points if know the pdf is the derivative of cdf.

note 2: If you lose 1 point if you didn't write the domain of the cdf.

6. (12%)

- (a) (5%) 若 $P_X(i) = \frac{m^i}{i!}e^{-m}$, $P_Y(j) = \frac{n^j}{j!}e^{-n}$, $i, j = 0, 1, 2, \dots$, $m, n > 0$, 且 X, Y 獨立。令 $Z = X + Y$, 證明 $P_Z(k) = \frac{(m+n)^k}{k!}e^{-(m+n)}$.
- (b) (7%) 已知早上和下午尖峰時間的車禍件數各自呈 Poisson 分配, 平均每個早上發生 1 件, 每個下午發生 2 件。若兩時段發生車禍數互相獨立, 求一天內兩尖峰時段總計發生 3 件以上車禍的機率。

Solution:

(a)

$$P_Z(k) = P(Z = k) = P(X + Y = k) = \sum_{i=0}^k P(X = i, Y = k - i) \quad (1)$$

Since X and Y are independent, we have

$$\begin{aligned} \sum_{i=0}^k P(X = i, Y = k - i) &= \sum_{i=0}^k P(X = i)P(Y = k - i) & (2) \\ &= \sum_{i=0}^k \left(\frac{m^i}{i!}e^{-m} \right) \cdot \left(\frac{n^{k-i}}{(k-i)!}e^{-n} \right) \\ &= e^{-(m+n)} \sum_{i=0}^k \frac{1}{i!(k-i)!} m^i \cdot n^{k-i} \\ &= \frac{e^{-(m+n)}}{k!} \sum_{i=0}^k \frac{k!}{i!(k-i)!} m^i \cdot n^{k-i} \\ &= \frac{e^{-(m+n)}}{k!} \sum_{i=0}^k C_i^k m^i \cdot n^{k-i} \\ &= \frac{e^{-(m+n)}}{k!} (m+n)^k \end{aligned}$$

- (b) The total number of accident in morning and evening is still a Poisson distribution, and its means is $1 + 2 = 3$.

Hence its probability function is $P(k) = \frac{3^k}{k!}e^{-3}$

The probability of no accident in morning and evening is $\frac{3^0}{0!}e^{-3} = e^{-3}$

The probability of one accident in morning and evening is $\frac{3^1}{1!}e^{-3} = 3e^{-3}$

The probability of two accident in morning and evening is $\frac{3^2}{2!}e^{-3} = \frac{9}{2}e^{-3}$

Therefore, the probability of having 3 or more accidents in morning and evening is $1 - e^{-3} - 3e^{-3} - \frac{9}{2}e^{-3} = 1 - \frac{17}{2}e^{-3}$

Grading Criteria

- (a) Write down (1) get 1 point

Use the independent property and write down (2) get 2 points

The rest of 2 points depend on your proof, you should explain each step you wrote down.

Note that Poisson distribution is a discrete probability distribution, you should use summation instead of integration!

- (b) 2 points for knowing the number of accident in morning and evening is still poisson distribution and its means is 3

Calculate the probability of 0, 1, 2 accident in morning and evening correctly get 1 point respectively

The rest of 2 point depend on your calculation of final answer.

7. (16%) 假設收銀員甲、乙服務一位客人的時間分別呈指數分配，平均時間各為 2 分鐘和 4 分鐘，並且甲、乙的服務是相互獨立的。

(a) (6%) 甲正開始服務一位客人，有另一人排隊。此時你排甲的櫃檯，等候時間的機率密度函數是 $f_X(t) = \frac{1}{4}te^{-\frac{1}{2}t}$, $t > 0$. 求等候時間的期望值。

(b) (10%) 承上題，乙正開始服務一位客人，沒有人排隊。求你由甲櫃檯換到乙櫃檯會提早接受到服務的機率。

Solution:

(a)

$$E(X) = \int_0^{\infty} t \cdot \frac{1}{4}te^{-\frac{1}{2}t} dt \quad (2\%)$$

$$= \left(-\frac{1}{2}t^2e^{-\frac{1}{2}t}\right)\Big|_0^{\infty} + \int_0^{\infty} te^{-\frac{1}{2}t} dt \quad (1\%)$$

$$= \left(-\frac{1}{2}t^2e^{-\frac{1}{2}t}\right)\Big|_0^{\infty} + \left(-2te^{-\frac{1}{2}t}\right)\Big|_0^{\infty} + \int_0^{\infty} 2e^{-\frac{1}{2}t} dt \quad (1\%)$$

$$= \left(-\frac{1}{2}t^2e^{-\frac{1}{2}t}\right)\Big|_0^{\infty} + \left(-2te^{-\frac{1}{2}t}\right)\Big|_0^{\infty} + \left(-4e^{-\frac{1}{2}t}\right)\Big|_0^{\infty}$$

$$= 0 + 0 + 4$$

$$= 4 \quad (2\%)$$

(b)

(2%) 排乙時等候時間 Y ，則

$$f_Y(y) = \frac{1}{4}e^{-\frac{y}{4}}$$

(2%) 聯合機率密度

$$f(x, y) = f_X(x) \cdot f_Y(y) = \frac{1}{16}xe^{-\frac{x}{2}}e^{-\frac{y}{4}}$$

(6%) 提早接受到服務的機率

$$\begin{aligned} P &= \iint_{0 \leq y \leq x} f(x, y) dx dy \\ &= \int_0^{\infty} \int_0^x \frac{1}{16}xe^{-\frac{x}{2}}e^{-\frac{y}{4}} dy dx \\ &= \int_0^{\infty} \left(-\frac{1}{4}xe^{-\frac{3x}{4}} + \frac{1}{4}xe^{-\frac{x}{2}}\right) dx \\ &= \frac{5}{9} \end{aligned}$$