

### 5.3.2 The Derived Distributions: Student's $t$ and Snedecor's $F$

Definition Let  $X_1, \dots, X_n$  be a random sample from a  $N(\mu, \sigma^2)$  distribution. The quantity  $(\bar{X} - \mu)/(S/\sqrt{n})$  has Student's  $t$  distribution with  $n - 1$  degrees of freedom. Equivalently, a random variable  $T$  has Student's  $t$  distribution with  $p$  degrees of freedom, and we write  $T \sim t_p$  if it has pdf

$$f_T(t) = \frac{\Gamma(\frac{p+1}{2})}{\Gamma(\frac{p}{2})} \frac{1}{(p\pi)^{1/2}} \frac{1}{(1 + t^2/p)^{(p+1)/2}}, \quad -\infty < t < \infty.$$

Notice that if  $p = 1$ , then  $f_T(t)$  becomes the pdf of the Cauchy distribution, which occurs for samples of size 2.

The derivation of the  $t$  pdf is straightforward. Let  $U \sim N(0, 1)$ , and  $V \sim \chi_p^2$ . If they are independent, the joint pdf is

$$f_{U,V}(u, v) = \frac{1}{\sqrt{2\pi}} e^{-u^2/2} \frac{1}{\Gamma(p/2)2^{p/2}} v^{\frac{p}{2}-1} e^{-v/2}, \quad -\infty < u < \infty, \quad 0 < v < \infty.$$

Make the transformation

$$t = \frac{u}{\sqrt{v/p}}, \quad w = v,$$

and integrate out  $w$ , we can get the marginal pdf of  $t$ .

Student's  $t$  has no mgf because it does not have moments of all orders. In fact, if there are  $p$  degrees of freedom, then there are only  $p - 1$  moments. It is easy to check that

$$ET_p = 0, \quad \text{if } p > 1,$$

$$\text{Var}T_p = \frac{p}{p-2}, \quad \text{if } p > 2.$$

Example Let  $X_1, \dots, X_n$  be a random sample from  $N(\mu_X, \sigma_X^2)$  population, and let  $Y_1, \dots, Y_m$  be a random sample from an independent  $N(\mu_Y, \sigma_Y^2)$  population. If we were interested in comparing the variability of the populations, one quantity of interest would be the ratio  $\sigma_X^2/\sigma_Y^2$ . Information about this ratio is contained in  $S_X^2/S_Y^2$ , the ratio of sample variances. The  $F$  distribution allows us to compare these quantities by giving us a distribution of

$$\frac{S_X^2/S_Y^2}{\sigma_X^2/\sigma_Y^2} = \frac{S_X^2/\sigma_X^2}{S_Y^2/\sigma_Y^2}.$$

Definition Let  $X_1, \dots, X_n$  be a random sample from  $N(\mu_X, \sigma_X^2)$  population, and let  $Y_1, \dots, Y_m$  be a random sample from an independent  $N(\mu_Y, \sigma_Y^2)$  population. The random variable  $F = \frac{S_X^2/\sigma_X^2}{S_Y^2/\sigma_Y^2}$  has Snedecor's F distribution with  $n - 1$  and  $m - 1$  degrees of freedom. Equivalently, the random variable  $F$  has the  $F$  distribution with  $p$  and  $q$  degrees of freedom if it has pdf

$$f_F(x) = \frac{\Gamma(\frac{p+q}{2})}{\Gamma(\frac{p}{2})\Gamma(\frac{q}{2})} \left(\frac{p}{q}\right)^{p/2} \frac{x^{(p/2)-1}}{[1 + (p/q)x]^{(p+q)/2}}, \quad 0 < x < \infty.$$

A variance ratio may have an  $F$  distribution even if the parent populations are not normal. Kelker (1970) has shown that as long as the parent populations have a certain type of symmetric, then the variance ratio will have an  $F$  distribution.

Example To see how the  $F$  distribution may be used for inference about the true ratio of population variances, consider the following. The quantity  $\frac{S_X^2/\sigma_X^2}{S_Y^2/\sigma_Y^2}$  has an  $F_{n-1, m-1}$  distribution. We can calculate

$$\begin{aligned} EF_{n-1, m-1} &= E\left(\frac{\chi_{n-1}^2/(n-1)}{\chi_{m-1}^2/(m-1)}\right) \\ &= E(\chi_{n-1}^2/(n-1))E((m-1)/(\chi_{m-1}^2)) = (m-1)/(m-3). \end{aligned}$$

Note this last expression is finite and positive only if  $m > 3$ . Removing expectations, we have for reasonably large  $m$ ,

$$\frac{S_X^2/S_Y^2}{\sigma_X^2/\sigma_Y^2} \approx \frac{m-1}{m-3} \approx 1,$$

as we might expect.

The  $F$  distribution has many interesting properties and is related to a number of other distributions.

Theorem 5.3.8

- a. If  $X \sim F_{p,q}$ , then  $1/X \sim F_{q,p}$ ; that is, the reciprocal of an  $F$  random variable is again an  $F$  random variable.
- b. If  $X \sim t_q$ , then  $X^2 \sim F_{1,q}$ .
- c. If  $X \sim F_{p,q}$ , then  $(p/q)X/(1 + (p/q)X) \sim \text{beta}(p/2, q/2)$ .