

國立臺灣大學數學系  
九十六學年度博士班入學考試試題  
科目：迴歸分析

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1. Consider the multiple linear regression model  $Y_i = \beta_0 + \beta_1 x_{i1} + \cdots + \beta_p x_{ip} + \epsilon_i$ ,  $i = 1, \dots, n$ , where  $\epsilon_i$ 's are uncorrelated with zero mean and constant variance  $\sigma^2$ . Let  $\hat{\beta} = (\hat{\beta}_0, \hat{\beta}_1, \dots, \hat{\beta}_p)^T$  be the least squares estimator of  $\beta = (\beta_0, \beta_1, \dots, \beta_p)^T$ .

(1a) (10%) Let  $e_{i(j)}$ 's and  $d_{ij}$ 's denote separately the residuals computed based on the regression models  $Y_i = \beta_{0j} + \sum_{l \neq j} \beta_{lj} x_{il} + \eta_i$  and  $x_{ij} = \gamma_0 + \sum_{l \neq j} \gamma_{lj} x_{il} + \xi_i$ ,  $i = 1, \dots, n$ ;  $j = 1, \dots, p$ . Show that  $Var(\hat{\beta}_j)$  increases as  $R_j^2 = 1 - (\sum_{i=1}^n d_{ij}^2) / (\sum_{i=1}^n e_{i(j)}^2)$  increases.

(1b) (10%) Let  $R^2$  be the coefficient of determination and  $r$  be the sample correlation coefficient of  $(Y_1, \hat{Y}_1), \dots, (Y_n, \hat{Y}_n)$ , where  $\hat{Y}_i = \hat{\beta}_0 + \hat{\beta}_1 x_{i1} + \cdots + \hat{\beta}_p x_{ip}$ ,  $i = 1, \dots, n$ . Show that  $r^2 = R^2$ .

(1c) (10%) Assume that  $\epsilon_1, \dots, \epsilon_n \stackrel{i.i.d.}{\sim} N(0, \sigma^2)$ . Derive the distribution of the studentized residual of the first observation  $r_1^* = e_1 / \sqrt{MSE_{(1)}(1 - h_{11})}$ , where  $e_1$  and  $h_{11}$  are separately the residual and the leverage point of the first case and  $MSE_{(1)}$  is the mean squared error obtained after deleting the first case.

2. Consider the linear model  $Y = X\beta + \epsilon$ , where  $\epsilon \sim N_n(0, \sigma^2 I_n)$  and  $X$  is a  $n \times p$  full rank matrix. Let  $\tilde{\beta}$  and  $e$  be separately the maximum likelihood estimator and the residual vector.

(2a) (10%) Show that  $\tilde{\beta}$  and  $SSE = e^T e$  are independent.

(2b) (7%) (8%) Construct a  $1 - \alpha$ ,  $0 < \alpha < 1$ , confidence interval for  $x_0^T \beta$  and a  $1 - \alpha$  prediction interval for the future independent run  $Y_0 = x_0^T \beta + \epsilon_0$  with  $\epsilon_0 \sim N(0, \sigma^2)$ .

3. (5%) (5%) Consider the growth curve model  $Y_i^{(t)} = X\beta + \epsilon_i^{(t)}$ ,  $i = 1, \dots, n$ , where  $X$  is a  $m \times p$  full rank matrix with  $p < m$  and  $\epsilon_1^{(t)}, \dots, \epsilon_n^{(t)} \stackrel{i.i.d.}{\sim} (0, \Omega^{(t)})$ . Find the estimated generalized least squares estimator of  $\beta$  and an unbiased estimator of  $\Omega^{(t)}$ .

4. (5%) (5%) (5%) Consider the model  $y_t = \beta_0 + \beta_1 x_{t1} + \cdots + \beta_p x_{tp} + \epsilon_t$ ,  $t = 1, \dots, n$ , where  $\epsilon_t = \rho \epsilon_{t-1} + u_t$  with  $u_t$ 's  $\stackrel{i.i.d.}{\sim} N(0, \sigma_u^2)$ . Find the consistent estimators of  $\rho$  and  $\sigma_u^2$ , and the estimated generalized least squares estimator of  $\beta = (\beta_0, \beta_1, \dots, \beta_p)$ .

5. Consider the centered and scaled regression model  $\mathbf{Y} = \tilde{\mathbf{Z}}_{(s)}\boldsymbol{\beta}_{(s)} + \boldsymbol{\varepsilon}$ , where  $\tilde{\mathbf{Z}}_{(s)} = (\mathbf{1}, \mathbf{Z}_{(s)})$  with  $\mathbf{Z}_{(s)}$  being a  $n \times p$  standardized covariate matrix,  $\boldsymbol{\beta}_{(s)} = \begin{pmatrix} \gamma_0 \\ \boldsymbol{\beta}_{(0)} \end{pmatrix}$ , and  $\boldsymbol{\varepsilon} \sim (0, \sigma^2 I_n)$ . Assume that an approximate linear relationship exists among the column vectors of  $\mathbf{Z}_{(s)}$ .

(5a) (10%) Show that  $(R^{-1})_{jj} = VIF_j$ ,  $j = 1, \dots, p$ , where  $R = (\mathbf{Z}_{(s)}^T \mathbf{Z}_{(s)})$  and  $VIF_j$  is the variance inflation factor associates with the  $j$ th covariate.

(5b) (10%) Let  $\hat{\boldsymbol{\beta}}_{(0)}$  and  $\hat{\boldsymbol{\beta}}_{(0)R}$  denote separately the least squares estimator and the ridge regression estimator of  $\boldsymbol{\beta}_{(0)}$ . Show that  $Var(\hat{\boldsymbol{\beta}}_{(0)}) - Var(\hat{\boldsymbol{\beta}}_{0R})$  is at least positive semi-definite.